

1060-710

# Mathematical and Statistical Methods for Astrophysics

## Problem Set 6

Assigned 2010 October 19  
Due 2010 October 26

**Show your work on all problems!** Be sure to give credit to any collaborators, or outside sources used in solving the problems. Note that if using an outside source to do a calculation, you should use it as a reference for the method, and actually carry out the calculation yourself; it's not sufficient to quote the results of a calculation contained in an outside source.

### 1 Upper Limits

Consider an experiment designed to measure an unknown physical quantity  $x$ , which returns a value  $y$  whose pdf is defined by the likelihood function

$$f(y|x) = \frac{e^{-(y-x)^2/2\sigma^2}}{\sigma\sqrt{2\pi}} \quad (1.1)$$

- a) Suppose the experiment has been performed and the result  $\hat{y}$  has been found. Calculate the frequentist upper limit  $x_{\text{UL}}^{\text{freq}}$  at confidence level  $\alpha$ , defined by

$$\int_{\hat{y}}^{\infty} f(y|x_{\text{UL}}^{\text{freq}}) dy = \alpha . \quad (1.2)$$

You should be able to write this with the help of the inverse complementary error function  $\text{erfc}^{-1}(\xi)$ . Note that  $\text{erfc}^{-1}(\xi)$  is positive if  $0 < \xi < 1$  and negative if  $1 < \xi < 2$ , and that  $\text{erfc}^{-1}(2 - \xi) = -\text{erfc}^{-1}(\xi)$

- b) Consider a Bayesian analysis with a uniform prior on  $x$ , so that by Bayes's theorem, the posterior is

$$f(x|y) = \frac{f(x)}{f(y)} f(y|x) = \mathcal{A} f(y|x) . \quad (1.3)$$

Using the explicit form of the likelihood (1.1) and the normalization requirement

$$\int_{-\infty}^{\infty} f(x|y) dx = 1 \quad (1.4)$$

find the value of  $\mathcal{A}$  and therefore the explicit form of the posterior  $f(x|y)$ .

- c) Supposing again that we've performed the experiment and found a result  $\hat{y}$ , find the Bayesian upper limit  $x_{\text{UL}}^{\text{Bayes}}$  at confidence level  $\alpha$ , defined by

$$\int_{-\infty}^{x_{\text{UL}}^{\text{Bayes}}} f(x|\hat{y}) dx = \alpha \quad (1.5)$$

- d) For the case where  $\alpha = 0.9$ , write  $x_{\text{UL}}^{\text{freq}}$  and  $x_{\text{UL}}^{\text{Bayes}}$  explicitly in terms of  $\hat{y}$  and  $\sigma$ , with any constants evaluated to three significant figures. (You'll need to refer to the explicit value of  $\text{erfc}^{-1}(\xi)$  for a particular  $\xi$ ; in matplotlib you can get access to the inverse complementary error function via `from scipy.special import erfcinv`.)
- e) Suppose now that  $x$  is physically constrained to be positive and let the prior be uniform for positive  $x$ , so that the posterior can be written in terms of the Heaviside step function

$$\Theta(x) = \begin{cases} 0 & x < 0 \\ 1 & x > 0 \end{cases} \quad (1.6)$$

as

$$f(x|y) = \frac{f(x)}{f(y)} f(y|x) = \mathcal{B} \Theta(x) f(y|x) . \quad (1.7)$$

Use the normalization condition

$$1 = \int_0^\infty f(x|y) dx = \mathcal{B} \int_0^\infty f(y|x) dx \quad (1.8)$$

to find the value of  $\mathcal{B}$  and therefore the explicit form of  $f(x|y)$ .

- f) Supposing again that we've performed the experiment and found a result  $\hat{y}$ , calculate the Bayesian upper limit  $x_{\text{UL}}^{\text{Bayes}+}$  associated with the posterior (1.7), defined by

$$\int_0^{x_{\text{UL}}^{\text{Bayes}+}} f(x|\hat{y}) dx = \alpha \quad (1.9)$$

## 2 Marginalization and the Inverse Fisher Matrix

Consider two variables  $X_1$  and  $X_2$  whose joint pdf is a Gaussian with zero mean:

$$f(\mathbf{x}) = \frac{\sqrt{\det \mathbf{F}}}{2\pi} \exp \left[ -\frac{1}{2} \mathbf{x}^T \mathbf{F} \mathbf{x} \right] = \frac{\sqrt{F_{11}F_{22} - F_{12}^2}}{2\pi} \exp \left[ -\frac{F_{11}}{2}(x_1)^2 - F_{12}x_1x_2 - \frac{F_{22}}{2}(x_2)^2 \right] \quad (2.1)$$

where  $\mathbf{F}$  is some symmetric, positive definite matrix.

- a) Show that  $\mathbf{F}$  is indeed the Fisher matrix.
- b) Marginalize over  $x_2$  and show that the resulting pdf for  $x_1$  is a Gaussian whose variance is the 1,1 component of the inverse Fisher matrix  $\mathbf{F}^{-1}$ :

$$f_{X_1}(x_1) = \int_{-\infty}^{\infty} f(x_1, x_2) dx_2 = \frac{1}{\sqrt{2\pi (F^{-1})_{11}}} \exp \left( -\frac{x_1^2}{2(F^{-1})_{11}} \right) \quad (2.2)$$

### 3 Least Squares and Chi-Squared

Consider measurements  $\{y_i\}$  taken at times  $\{t_i\} = \{-1, 0, 1, 2\}$ . We wish to fit these measurements with a straight-line model with predicted expectation values  $\mu_i = \lambda_1 + \lambda_2 t_i$ . The model predicts measurements which differ from  $\mu_i$  by uncorrelated Gaussian errors with standard deviations  $\{\sigma_i\} = \{\sqrt{2}, 1, \sqrt{2}, \sqrt{3}\}$ .

- a) Find the matrix  $\mathbf{A}$  describing the linear relationship  $\boldsymbol{\mu} = \mathbf{A}\boldsymbol{\lambda}$ , i.e.,

$$\begin{pmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \mu_4 \end{pmatrix} = \mathbf{A} \begin{pmatrix} \lambda_1 \\ \lambda_2 \end{pmatrix} \quad (3.1)$$

- b) Since the errors are uncorrelated, the standard deviations are described by a matrix

$$\boldsymbol{\sigma} = \begin{pmatrix} \sigma_1 & 0 & 0 & 0 \\ 0 & \sigma_2 & 0 & 0 \\ 0 & 0 & \sigma_3 & 0 \\ 0 & 0 & 0 & \sigma_4 \end{pmatrix}. \quad (3.2)$$

Construct the matrix  $\mathbf{A}^T \boldsymbol{\sigma}^{-2} \mathbf{A}$  and find its inverse  $[\mathbf{A}^T \boldsymbol{\sigma}^{-2} \mathbf{A}]^{-1}$ . (Since it is a  $2 \times 2$  matrix, you should actually be able to invert it by hand.)

- c) In class we showed that if the measured values are  $\mathbf{y}$ , the maximum likelihood estimates of the parameters will be  $\hat{\boldsymbol{\lambda}}(\mathbf{y}) = [\mathbf{A}^T \boldsymbol{\sigma}^{-2} \mathbf{A}]^{-1} \mathbf{A}^T \boldsymbol{\sigma}^{-2} \mathbf{y}$ . Work out the elements of the matrix appearing for this problem in

$$\begin{pmatrix} \hat{\lambda}_1 \\ \hat{\lambda}_2 \end{pmatrix} = [\mathbf{A}^T \boldsymbol{\sigma}^{-2} \mathbf{A}]^{-1} \mathbf{A}^T \boldsymbol{\sigma}^{-2} \begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{pmatrix} \quad (3.3)$$

- d) Suppose we measure  $\{y_i\} = \{1.07241020, 0.40438919, 2.89906726, 8.98526374\}$ . Calculate, to three significant figures,

- i) The best-fit parameters  $\hat{\lambda}_1$  and  $\hat{\lambda}_2$
- ii) The  $\chi^2$  value relating the data to the best-fit model,

$$\chi^2 = (\mathbf{y} - \mathbf{A}\hat{\boldsymbol{\lambda}})^T \boldsymbol{\sigma}^{-2} (\mathbf{y} - \mathbf{A}\hat{\boldsymbol{\lambda}}) \quad (3.4)$$

- iii) The  $p$  value, i.e., probability that data generated according to the model would have a  $\chi^2$  equal to or higher than the one observed.